A discrete-time Matsumoto-Yor theorem

Charlie Hérent

Université Paris Cité (MAP5)

CORTIPOM conference, Le Croisic 10 juin 2025

Background and literature

- Pitman's theorem
- Matsumoto-Yor's theorem and extension

2 A discrete-time Matsumoto-Yor theorem

- Random walk on a subgroup of SL₂
- Main theorems obtained

Charlie Hérent (MAP5)

Pitman's theorem (1975)

• Let $(B_t)_{t\geq 0}$ be a one-dimensional Brownian motion, then

$$\left(\mathcal{Z}_t := B_t - 2\inf_{0 \le s \le t} B_s\right)_{t \ge 0}$$

is a Bessel process of dimension 3 starting from 0.

Charlie Hérent (MAP5)

Pitman's theorem (1975)

• Let $(B_t)_{t\geq 0}$ be a one-dimensional Brownian motion, then

$$\left(\mathcal{Z}_t := B_t - 2\inf_{0 \le s \le t} B_s\right)_{t \ge 0}$$

is a Bessel process of dimension 3 starting from 0.

• The law of B_t conditionally to $(\mathcal{Z}_s)_{0 \le s \le t}$ is uniform on $[-\mathcal{Z}_t, \mathcal{Z}_t]$.

Pitman's theorem (1975)

• Let $(B_t)_{t\geq 0}$ be a one-dimensional Brownian motion, then

$$\left(\mathcal{Z}_t := B_t - 2\inf_{0 \le s \le t} B_s\right)_{t \ge 0}$$

is a Bessel process of dimension 3 starting from 0.

• The law of B_t conditionally to $(\mathcal{Z}_s)_{0 \le s \le t}$ is uniform on $[-\mathcal{Z}_t, \mathcal{Z}_t]$.

A Bessel process of dimension 3 is a stochastic process $(\mathcal{Z}_t)_{t\geq 0}$ such that $(\mathcal{Z}_t)_{t\geq 0} \stackrel{\text{law}}{=} ||B_t^{(3)}||$ with $B^{(3)}$ a Brownian motion of dimension 3.

Pitman's theorem (1975)

• Let $(B_t)_{t\geq 0}$ be a one-dimensional Brownian motion, then

$$\left(\mathcal{Z}_t := B_t - 2\inf_{0 \le s \le t} B_s\right)_{t \ge 0}$$

is a Bessel process of dimension 3 starting from 0.

• The law of B_t conditionally to $(\mathcal{Z}_s)_{0 \le s \le t}$ is uniform on $[-\mathcal{Z}_t, \mathcal{Z}_t]$.

A Bessel process of dimension 3 is a stochastic process $(\mathcal{Z}_t)_{t\geq 0}$ such that $(\mathcal{Z}_t)_{t\geq 0} \stackrel{\text{law}}{=} ||B_t^{(3)}||$ with $B^{(3)}$ a Brownian motion of dimension 3. A Bessel process of dimension 3 = "Brownian motion conditioned to stay forever positive" using Doob's transform.

Matsumoto-Yor's theorem

Charlie Hérent (MAP5)

Let $(B_t^{(\mu)})_{t\geq 0}$ be a one-dimensional Brownian motion with drift $\mu \in \mathbb{R}$.

Matsumoto-Yor's theorem

Let $(B_t^{(\mu)})_{t\geq 0}$ be a one-dimensional Brownian motion with drift $\mu \in \mathbb{R}$.

Matsumoto-Yor's theorem (2000)

$$Z_t^{(\mu)} := e^{B_t^{(\mu)}} \int_0^t e^{-2B_s^{(\mu)}} ds$$

is a diffusion process on \mathbb{R}_+ with infinitesimal generator given by :

$$\frac{1}{2}z^{2}\frac{d^{2}}{dz^{2}} + \left[\left(\frac{1}{2} + \mu\right)z + \left(\frac{K_{1-\mu}}{K_{\mu}}\right)\left(\frac{1}{z}\right)\right]\frac{d}{dz}$$

where K_{μ} is a Macdonald function (modified Bessel function of the second kind).

H. Matsumoto and M. Yor. An analogue of Pitman's 2M - X theorem for exponential Wiener functionals. I. A time-inversion approach. Nagoya Math. J. 159 125 - 166, 2000.

Charlie Hérent (MAP5)

Let $A \subset \mathbb{R}^d$ and $\varphi : \mathbb{R}^d \to \mathbb{R}$ be a continuous function such that :

$$\int_A e^{-\varphi(x)} dx < +\infty.$$

Then :

$$\lim_{\theta \to +\infty} \frac{1}{\theta} \log \left(\int_{A} e^{-\theta \varphi(x)} dx \right) = - \inf_{x \in A} \varphi(x).$$

The Matsumoto–Yor theorem can be interpreted through $(B_t)_{t\geq 0}$:

$$\mathbf{B}_{t} = \begin{pmatrix} e^{B_{t}^{(\mu)}} & 0\\ e^{B_{t}^{(\mu)}} \int_{0}^{t} e^{-2B_{s}^{(\mu)}} ds & e^{-B_{t}^{(\mu)}} \end{pmatrix}$$

The Matsumoto–Yor theorem can be interpreted through $(B_t)_{t\geq 0}$:

$$\mathbf{B}_{t} = \begin{pmatrix} e^{B_{t}^{(\mu)}} & 0\\ e^{B_{t}^{(\mu)}} \int_{0}^{t} e^{-2B_{s}^{(\mu)}} ds & e^{-B_{t}^{(\mu)}} \end{pmatrix}$$

This process is the solution to the following SDE in the sense of Stratonovitch :

Charlie Hérent (MAP5)

The Matsumoto–Yor theorem can be interpreted through $(B_t)_{t\geq 0}$:

$$\mathbf{B}_{t} = \begin{pmatrix} e^{B_{t}^{(\mu)}} & 0\\ e^{B_{t}^{(\mu)}} \int_{0}^{t} e^{-2B_{s}^{(\mu)}} ds & e^{-B_{t}^{(\mu)}} \end{pmatrix}$$

This process is the solution to the following SDE in the sense of Stratonovitch :

$$d\mathbf{B}_t = \mathbf{B}_t \circ \begin{pmatrix} dB_t^{(\mu)} & 0 \\ dt & -dB_t^{(\mu)} \end{pmatrix}$$
 with $\mathbf{B}_0 = I_2$.

The Matsumoto–Yor theorem can be interpreted through $(B_t)_{t\geq 0}$:

$$\mathbf{B}_{t} = \begin{pmatrix} e^{B_{t}^{(\mu)}} & 0\\ e^{B_{t}^{(\mu)}} \int_{0}^{t} e^{-2B_{s}^{(\mu)}} ds & e^{-B_{t}^{(\mu)}} \end{pmatrix}$$

This process is the solution to the following SDE in the sense of Stratonovitch :

$$d\mathbf{B}_t = \mathbf{B}_t \circ \begin{pmatrix} dB_t^{(\mu)} & 0 \\ dt & -dB_t^{(\mu)} \end{pmatrix}$$
 with $\mathbf{B}_0 = I_2$.

This SDE, in its general form, was introduced by Biane, Bougerol and O'Connell.

The Matsumoto–Yor theorem can be interpreted through $(B_t)_{t\geq 0}$:

$$\mathbf{B}_{t} = \begin{pmatrix} e^{B_{t}^{(\mu)}} & 0\\ e^{B_{t}^{(\mu)}} \int_{0}^{t} e^{-2B_{s}^{(\mu)}} ds & e^{-B_{t}^{(\mu)}} \end{pmatrix}$$

This process is the solution to the following SDE in the sense of Stratonovitch :

$$d\mathbf{B}_t = \mathbf{B}_t \circ \begin{pmatrix} dB_t^{(\mu)} & 0 \\ dt & -dB_t^{(\mu)} \end{pmatrix}$$
 with $\mathbf{B}_0 = I_2$.

This SDE, in its general form, was introduced by Biane, Bougerol and O'Connell.

O'Connell and then Chhaibi extended the Matsumoto-Yor theorem in the context of semisimple Lie groups.

C. Hérent. *A discrete-time Matsumoto–Yor theorem*. Preprint 2024. (hal-04683351), arXiv :2409.01044. Accepted for publication in ESAIM : Probability and Statistics.

Charlie Hérent (MAP5)

C. Hérent. *A discrete-time Matsumoto–Yor theorem*. Preprint 2024. (hal-04683351), arXiv :2409.01044. Accepted for publication in ESAIM : Probability and Statistics.

Let $\gamma := (\gamma_n)_{n \in \mathbb{N}}$ be a family of i.i.d. random variables and let $\delta \in \mathbb{R}^*$ be a deterministic parameter. Let $(X_n)_{n \in \mathbb{N}}$, $(Z_n)_{n \in \mathbb{N}}$ be two processes such as :

C. Hérent. *A discrete-time Matsumoto–Yor theorem*. Preprint 2024. (hal-04683351), arXiv :2409.01044. Accepted for publication in ESAIM : Probability and Statistics.

Let $\gamma := (\gamma_n)_{n \in \mathbb{N}}$ be a family of i.i.d. random variables and let $\delta \in \mathbb{R}^*$ be a deterministic parameter. Let $(X_n)_{n \in \mathbb{N}}$, $(Z_n)_{n \in \mathbb{N}}$ be two processes such as :

$$\begin{pmatrix} X_{n+1} & 0 \\ Z_{n+1} & X_{n+1}^{-1} \end{pmatrix} = \begin{pmatrix} X_n & 0 \\ Z_n & X_n^{-1} \end{pmatrix} \begin{pmatrix} \gamma_n & 0 \\ \delta & \gamma_n^{-1} \end{pmatrix}$$

with,

$$\begin{pmatrix} X_0 & 0 \\ Z_0 & X_0^{-1} \end{pmatrix} = I_2.$$

By recurrence, we have the formula for $n \in \mathbb{N}$:

$$\begin{pmatrix} X_n & 0 \\ Z_n & X_n^{-1} \end{pmatrix} = \begin{pmatrix} \prod_{i=0}^{n-1} \gamma_i & 0 \\ \delta \sum_{k=0}^{n-1} \prod_{i=0}^{k-1} \gamma_i^{-1} \prod_{j=k+1}^{n-1} \gamma_j & \prod_{i=0}^{n-1} \gamma_i^{-1} \end{pmatrix}.$$

By recurrence, we have the formula for $n \in \mathbb{N}$:

$$\begin{pmatrix} X_n & 0 \\ Z_n & X_n^{-1} \end{pmatrix} = \begin{pmatrix} \prod_{i=0}^{n-1} \gamma_i & 0 \\ \delta \sum_{k=0}^{n-1} \prod_{i=0}^{k-1} \gamma_i^{-1} \prod_{j=k+1}^{n-1} \gamma_j & \prod_{i=0}^{n-1} \gamma_i^{-1} \end{pmatrix}.$$

In the following, we will assume (except for the continuous limit of the random walk) that $\delta = 1$ to simplify certain expressions.

For the common law of γ_i , we take a $GIG(\lambda, a, a)$ distribution.

Charlie Hérent (MAP5)

For the common law of γ_i , we take a $GIG(\lambda, a, a)$ distribution. The probability density function for this distribution is given by :

$$f(x) := \frac{1}{2K_{\lambda}(a^2)} x^{\lambda-1} e^{-\frac{a^2}{2}(x+\frac{1}{x})} \mathbb{1}_{x>0}$$

where a > 0, $\lambda \in \mathbb{R}$ and K_{λ} is the Macdonald function.

Theorem (Markov property of the process Z)

Charlie Hérent (MAP5)

• The process $(Z_n)_{n \in \mathbb{N}^*}$ is a homogeneous Markov chain from $Z_1 = 1$ with transition kernel given, for x > 0, by :

$$Q(x, dy) = \left(\frac{1}{2K_{\lambda}(a^2)}\right) \frac{K_{\lambda}\left(\frac{a^2}{y}\right)}{K_{\lambda}\left(\frac{a^2}{x}\right)} \frac{1}{y} e^{-\frac{a^2(x^2+y^2+1)}{2xy}} \mathbb{1}_{\mathbb{R}^{*}_{+}}(y) dy.$$

Theorem (Markov property of the process Z)

• The process $(Z_n)_{n \in \mathbb{N}^*}$ is a homogeneous Markov chain from $Z_1 = 1$ with transition kernel given, for x > 0, by :

$$Q(x, dy) = \left(\frac{1}{2K_{\lambda}(a^2)}\right) \frac{K_{\lambda}\left(\frac{a^2}{y}\right)}{K_{\lambda}\left(\frac{a^2}{x}\right)} \frac{1}{y} e^{-\frac{a^2(x^2+y^2+1)}{2xy}} \mathbb{1}_{\mathbb{R}^{*}_{+}}(y) dy.$$

ullet In addition, for all $n\in\mathbb{N}^*$:

$$\mathbb{P}(X_n \in dx | Z_n, \cdots, Z_1) = \Lambda(Z_n, dx)$$
 a.s.

with, for z > 0: $\Lambda(z, dx) = \frac{1}{2\kappa_{\lambda}\left(\frac{a^2}{z}\right)} x^{\lambda-1} e^{-\frac{a^2}{2z}\left(x+\frac{1}{x}\right)} \mathbb{1}_{\mathbb{R}^{*}_{+}}(x) dx.$

Theorem

Let $(\gamma_n)_{n \in \mathbb{N}}$ be a sequence of i.i.d. random variables with a density function \mathcal{C}^1 supported on \mathbb{R}_+ . Then, there exists a family of probability measures $(\Lambda(z, \cdot))_{z>0}$ such that, for all $n \in \mathbb{N}^*$:

$$\mathbb{P}(X_n \in dx | Z_n, \cdots, Z_1) = \Lambda(Z_n, dx)$$
 p.s.

if and only if the sequence $(\gamma_n)_{n \in \mathbb{N}}$ is distributed with a $GIG(\lambda, a, a)$ law with parameters a > 0 and $\lambda \in \mathbb{R}$.

Charlie Hérent (MAP5)

Dufresne's identity establishes the following equality in law for $\mu > 0$:

Dufresne's identity establishes the following equality in law for $\mu > 0$:

$$\int_0^{+\infty} e^{-2B_s^{(\mu)}} ds \stackrel{\text{law}}{=} \frac{1}{2\xi}$$

where ξ is a random variable with distribution $Gamma(\mu)$, that is, with probability density $\frac{1}{\Gamma(\mu)}x^{\mu-1}e^{-x}\mathbb{1}_{x>0}$.

12 / 15

Theorem

For $\lambda > 0$, the distribution of the random variable

$$N_{\infty}^{(\lambda)} = \sum_{k=0}^{+\infty} \gamma_k^{-1} \left(\prod_{i=0}^{k-1} \gamma_i^{-1}\right)^2$$

is given by the inverse-gamma density distribution π defined by :

$$\pi(x) = \frac{a^{2\lambda}}{2^{\lambda} \Gamma(\lambda)} x^{-\lambda-1} e^{-\frac{a^2}{2x}} \mathbb{1}_{x>0}$$

where Γ is Euler's gamma function.

Convergence of discrete processes to their continuous analogues

We have the following convergence for the random walk on the group with $\delta = \delta_n$ such that $\delta_n := \frac{1}{n}$ et $\gamma_i^{(\sqrt{n})} \sim GIG(\lambda, \sqrt{n}, \sqrt{n})$ when $n \to +\infty$:

$$\begin{pmatrix} X_{\lfloor nt \rfloor} & 0 \\ Z_{\lfloor nt \rfloor} & X_{\lfloor nt \rfloor}^{-1} \end{pmatrix} \xrightarrow{\text{law}} \begin{pmatrix} e^{B_t^{(\lambda)}} & 0 \\ e^{B_t^{(\lambda)}} \int_0^t e^{-2B_s^{(\lambda)}} ds & e^{-B_t^{(\lambda)}} \end{pmatrix}$$

14 / 15

Thank you for your attention !

æ

$$X_2 = \gamma_0 \gamma_1, \ X_3 = \gamma_0 \gamma_1 \gamma_2, \ Z_2 = \gamma_0^{-1} + \gamma_1, \ Z_3 = \gamma_0^{-1} \gamma_1^{-1} + \gamma_0^{-1} \gamma_2 + \gamma_1 \gamma_2.$$

$$X_2 = \gamma_0 \gamma_1, \ X_3 = \gamma_0 \gamma_1 \gamma_2, \ Z_2 = \gamma_0^{-1} + \gamma_1, \ Z_3 = \gamma_0^{-1} \gamma_1^{-1} + \gamma_0^{-1} \gamma_2 + \gamma_1 \gamma_2.$$

Theorem (Characterisation of GIG laws)

Let $\gamma_0, \gamma_1, \gamma_2$ be three i.i.d. random variables with a density function C^1 supported on \mathbb{R}_+ . If the following conditional laws are equal, for z, u > 0:

$$\mathcal{L}(X_3|Z_3=z,Z_2=u)=\mathcal{L}(X_2|Z_2=z),$$

then, γ_0 is distributed according to a law $GIG(\lambda, a, a)$ with a > 0 and $\lambda \in \mathbb{R}$.

Let *P* and *Q* be two Markov kernels on the measurable spaces (E, \mathcal{E}) and (F, \mathcal{F}) respectively. Consider a Markov kernel Λ from *F* to *E*, that is, an application :

$$\Lambda : (u, A) \mapsto \Lambda(u, A)$$
 with $u \in F$ and $A \in \mathcal{E}$

such that for all $u \in F$, $\Lambda(u, \cdot)$ is a probability measure on E, and for all $A \in \mathcal{E}$, $\Lambda(\cdot, A)$ belongs to the space of bounded measurable functions on F.

Definition (Intertwining)

We will then say that the Markov kernels P and Q are intertwined by Λ if we have the relation :

$$\Lambda P = Q\Lambda$$

where the kernel composition is defined by :

$$\Lambda P(u, dv) := \int_E P(y, dv) \Lambda(u, dy).$$

We will say that two semigroups of Markovian kernels are intertwined if this relationship occurs, with the same Λ , for all times.

Rogers-Pitman criterion

Let $\phi : E \to F$ be a measurable function, we define the kernel Φ such that : $\Phi f := f \circ \phi$. Let Λ be a Markov kernel from F to E verifying the two conditions :

(a) ΛΦ = I, the identity kernel on F.
(b) For all t ≥ 0, the Markov kernel from F to F, Q_t := ΛP_tΦ satisfies the identity : ΛP_t = Q_tΛ.

Let X be a Markov process with semigroup $(P_t)_{t\geq 0}$ and initial distribution $\lambda := \Lambda(u, \cdot)$ where $u \in F$. Then, for all $t \geq 0$ and $A \in \mathcal{E}$:

$$\mathbb{P}\left(X_t \in A | \phi \circ X_s, \ 0 \le s \le t\right) = \Lambda(\phi \circ X_t, A)$$
 a.s.

Moreover, $\phi \circ X$ is Markov with initial state u and semigroup $(Q_t)_{t \geq 0}$.